EN.601.774 Theory of Replicable ML

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Lecture 9

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https://www.adaptivedataanalysis.com

Domain $\mathcal{X} = \{0, 1\}^d, \, \mathcal{Y} = \{0, 1\}.$

Overfitting with "natural" adaptive SQs

Algorithm 1 Query learner

Inputs/Parameters: Sample $S \sim D^m$

1: $P = \emptyset$

2: for $i \in [d]$ do

 $\phi_i(x,y) = \begin{cases} 1, & x_i = y \\ 0, & o.w. \end{cases}$

 $a_i \leftarrow \frac{1}{m} \sum_{(x,y) \in S} [\phi(x,y)]$ if $a_i \ge \frac{1}{2} + \frac{1}{\sqrt{m}}$ then $P = P \cup i$

6:

end if 7:

return $f(x) = \lfloor \frac{1}{|P|} \sum_{i \in P} x_i \rceil$

9: end for

Claim 0.1. When D is the uniform distribution over $\mathcal{X} \times \mathcal{Y}$, \exists constant c such that with probability at least $1 - \delta$, if $d \ge c \max\{m, \log(1/\delta)\}$:

$$|acc_S(f) - acc_D(f)| \ge .49$$

Compare to the accuracy guarantee we have for non-adaptive statistical queries, from which we would expect

$$|acc_S(f) - acc_D(f)| \in O\left(\sqrt{\frac{\log(d/\delta)}{m}}\right).$$

Proof. Let

$$X_i = \begin{cases} 1, & i \in P \\ 0, & o.w. \end{cases}$$

Then

$$\Pr_{S \sim D^m}[X_i = 1] = \Pr_{S \sim D^m}\left[\frac{1}{m} \sum_{(x,y) \in S} \mathbb{1}[x_i = y] \ge \frac{1}{2} + \frac{1}{\sqrt{m}}\right]$$

Let $A_i = \frac{1}{m} \sum_{(x,y) \in S} \mathbb{1}[x_i = y]$. A_i is a binomial random variable with $\mathbb{E}_D[A_i] = \frac{1}{2}$ and standard deviation $\frac{1}{2\sqrt{m}}$, and therefore

$$\Pr_D[X_i = 1] = \Pr_D[A_i \ge \frac{1}{2} + \frac{1}{\sqrt{m}}] \in \Omega(1).$$

Therefore, each i gets added to P with constant probability. It follows that

$$\mathbb{E}_{S \sim D^m}[|P| = \sum_{i=1}^d X_i] = \Omega(d).$$

Recall what the Chernoff-Hoeffding inequality gives us for a sum of bounded r.v.'s $X_i \in [a_i, b_i]$, letting $S_d = \sum_{i=1}^d X_i$:

$$\Pr_{X_1,...X_d} [S_d \le \mathbb{E}[S_d] - t] \le e^{\frac{-2t^2}{\sum_{i=1}^m (b_i - a_i)^2}}$$

applied to |P|, there is a $t \in \Omega(d)$ such that we have

$$\begin{split} \Pr_{S \sim D^m}[|P| \in o(d)] &\leq \Pr_{X_1, \dots, X_d}[|P| \leq \mathbb{E}[|P|] - t] \\ &\leq e^{\frac{-2t^2}{\sum_{i=1}^d (b_i - a_i)^2}} \\ &= e^{\frac{-2t^2}{d}} \\ &\in e^{-\Omega(d)} \end{split}$$

So there exists some constant c_1 such that so long as $d > c_1 \log(1/\delta)$,

$$\Pr_{S \sim D^m}[|P| \in \Omega(d)] > 1 - \delta$$

Now let's see how this causes us to get an unreliable empirical estimate of $acc_S(f)$ if we reuse the sample S. Let $(x,y) \sim S$ be chosen uniformly at random. f(x) = y iff $\sum_{i \in P} \mathbb{1}[x_i = y] \geq \frac{|P|}{2}$. We have that for each $i \in P$, $\Pr[x_i = y] \geq \frac{1}{2} + \frac{1}{\sqrt{m}}$, and so

$$\mathbb{E}_{(x,y)\sim S}[\sum_{i\in P} \mathbb{1}[x_i = y]] \ge \frac{|P|}{2} + \frac{|P|}{\sqrt{m}}$$

Therefore f(x) = y unless $\sum_{i \in P} \mathbb{1}[x_i = y]$ is less than its expectation by at least $\frac{|P|}{\sqrt{m}}$. Applying Chernoff-Hoeffding to the sum of random variables $C = \sum_{i \in P} \mathbb{1}[x_i = y]$, we have

$$1 - acc_S(f) = \Pr_{(x,y) \sim S} [f(x) \neq y]$$

$$= \Pr_{(x,y) \sim S} [C \leq \mathbb{E}[X] - \frac{|P|}{\sqrt{m}}]$$

$$\leq e^{\frac{-2|P|^2}{m|P|}}$$

$$= e^{\frac{-2|P|}{m}}$$

So if $|P| > \frac{\ln(100)m}{2}$, $acc_S(f) = .99$. However, $acc_D(f) = 1/2$. We already showed that there exists c_1 such that $|P| \in \Omega(d)$ except with probability δ , so long as $d > c_1 \log(1/\delta)$. Therefore there exists c_2 such that so long as $d > c_2 m$, $|P| > \frac{\ln(100)m}{2}$, and $acc_S(f) = .99$. Letting $c = \max\{c_1, c_2\}$, it follows that there exists a c such that with probability at least $1 - \delta$, if $d \ge c \max\{m, \log(1/\delta)\}$:

$$|acc_S(f) - acc_D(f)| > .49$$

Observations

- The argument above still goes through when we don't answer queries with an exact empirical estimate, but instead add some noise on the order $o(\frac{1}{\sqrt{n}})$ to the estimate.
- We could have done a similar analysis using only the first k-1 of d features. Redoing the argument using only k statistical queries instead of d+1 gives a bound of

$$|acc_S(f) - acc_D(f)| \in \Omega(\sqrt{km}).$$

So the best confidence interval we can hope for with k adaptive statistical queries, answered by empirical estimate over reused data, is $O(\sqrt{\frac{k}{m}})$. Recall that for k non-adaptive statistical queries, our bound was $O(\sqrt{\frac{\log k}{m}})$.

- If we want a confidence interval of ε , reusing data in this way doesn't save us anything, since we would need $m \in \Omega(\frac{k}{\varepsilon^2})$ samples... which is k times what we would need for a single statistical query.
- Could we have made our adaptive algorithm non-adaptive? The first d queries will non-adaptive, so what if we just committed to estimating the error of every possible f we might have constructed in our algorithm, rather than the one that was chosen after looking at the results of our d non-adaptive queries. Would this give a better bound? There are 2^d many subsets of d variables that could be included in P, and therefore 2^d different functions f. So we would need to make $O(2^d)$ statistical queries, giving a bound of $O(\sqrt{\frac{\log 2^d}{m}}) = O(\sqrt{\frac{d}{m}})$, so no better than the adaptive version.
- Do replicable SQs help? Since the first d queries are non-adaptive, we know that so long as we use a large enough sample, we can guarantee

$$\Pr_{S_1, S_2, r}[f_{S_1}^r = f_{S_2}^r] > 1 - \rho$$

where $f_{S_i}^r = A(S_i; r)$. It follows that

$$\begin{split} \Pr_{S_{1},r}[acc_{S_{1}}(f^{r}_{S_{1}}) \geq \frac{1}{2} + \tau] &= \Pr_{S_{1},S_{2},r}[acc_{S_{1}}(f^{r}_{S_{2}}) \geq \frac{1}{2} + \tau \mid f^{r}_{S_{2}} = f^{r}_{S_{1}}] \cdot \Pr_{S_{1},S_{2},r}[f^{r}_{S_{2}} = f^{r}_{S_{1}}] \\ &+ \Pr_{S_{1},S_{2},r}[acc_{S_{1}}(f^{r}_{S_{1}}) \geq \frac{1}{2} + \tau \mid f^{r}_{S_{2}} = f^{r}_{S_{1}}] \cdot \Pr_{S_{1},S_{2},r}[f^{r}_{S_{2}} \neq f^{r}_{S_{1}}] \\ &\leq \Pr_{S_{1},S_{2},r}[acc_{S_{1}}(f^{r}_{S_{2}}) \geq \frac{1}{2} + \tau \mid f^{r}_{S_{2}} = f^{r}_{S_{1}}] \cdot \Pr_{S_{1},S_{2},r}[f^{r}_{S_{2}} = f^{r}_{S_{1}}] + \rho \\ &\leq \Pr_{S_{1},S_{2},r}[acc_{S_{1}}(f^{r}_{S_{2}}) \geq \frac{1}{2} + \tau] + \rho \\ &= \Pr_{S_{1},S_{2},r}[acc_{S_{1}}(f^{r}_{S_{2}}) - \mathbb{E}_{S_{1},S_{2},r}[acc_{S_{1}}(f^{r}_{S_{2}})] \geq \tau] + \rho \\ &\leq e^{-2\tau^{2}m} + \rho \\ &\in O(\rho) \end{split}$$

so long as we take $m \in \Omega(\frac{\log 1/\rho}{\tau^2})$.

However, to ensure that $\Pr_{S_1,S_2,r}[f_{S_2}^r \neq f_{S_1}^r] \leq \rho$, we need to make d non-adaptive replicable statistical queries with $\rho' = \rho/d$, so we need $O(\frac{d^2}{\tau^2\rho^2})$ samples. Which is already worse than resampling!